

Regular Paper

Estimating Membrane Resistance over Dendrite Using Markov Random Field

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Abstract: With developments in optical imaging over the past decade, statistical methods for estimating dendritic membrane resistance from observed noisy signals have been proposed. In most of previous studies, membrane resistance over a dendritic tree was assumed to be constant, or membrane resistance at a point rather than that over a dendrite was investigated. Membrane resistance, however, is actually not constant over a dendrite. In a previous study, a method was proposed in which membrane resistance value is expressed as a non-constant function of position on dendrite, and parameters of the function are estimated. Although this method is effective, it is applicable only when the appropriate function is known. We propose a statistical method, which does not express membrane resistance as a function of position on dendrite, for estimating membrane resistance over a dendrite from observed membrane potentials. We use the Markov random field (MRF) as a prior distribution of the membrane resistance. In the MRF, membrane resistance is not expressed as a function of position on dendrite, but is assumed to be smoothly varying along a dendrite. We apply our method to synthetic data to evaluate its efficacy, and show that even when we do not know the appropriate function, our method can accurately estimate the membrane resistance.

Keywords: dendrite, membrane resistance, Markov random field, cable equation, membrane potential imaging

1. Introduction

Information processing in neural systems is suggested to be dependent on how the membrane properties are varying along dendritic trees [1], [2], [3], [4], [5], [6], [7]. In hippocampal CA1 pyramidal neurons, for example, the membrane resistance varies sigmoidally along a dendritic tree. A recent computational study showed that this sigmoidicity improves the efficiency of information propagation from the distal to proximal parts [6].

With developments in optical imaging over the past decade, several statistical methods for estimating membrane properties, especially membrane resistance, from fluorescence intensity have been proposed [4], [5], [7], [8], [9]. Optical imaging, however, has a low signal-to-noise ratio [10], [11], [12], [13], [14], [15], [16], [17], so accurately estimating membrane resistance over a dendritic tree is challenging. In previous studies, membrane resistance over a dendritic tree was assumed to be constant, or membrane resistance at a point rather than that over a dendrite was investigated. We previously proposed a method in which membrane resistance value is expressed as a non-constant function of position on dendrite, and parameters of the function are estimated [4], [5], [7]. Although this method can accurately es-

timate membrane resistance over a dendrite, it is applicable only when we know the appropriate function. Thus, developing methods for estimating membrane resistance over a dendrite remains a challenge.

For this study, we propose a statistical method, which does not express membrane resistance as a function of position on dendrite, for estimating membrane resistance over dendrite from observed noisy signals. For this purpose, we use the Markov random field (MRF) [18], [19] as a prior distribution of the membrane-resistance. In the MRF, membrane resistance is not expressed as a function of position on dendrite, but is assumed to be smoothly varying along a dendrite. This smoothness prior expresses a physiological premise that spatially adjacent membrane resistances take similar values. Additionally, the dynamics of membrane potential corresponding to a state in dendritic systems is expressed using the cable equation [20], [21], and the observation process is expressed using a Gaussian process. We estimate parameters, namely, membrane-resistance over a dendrite by using the expectation-maximization (EM) algorithm [22]. We applied our method to synthetic data to evaluate its efficacy, and show that even when we do not know the appropriate function, our method can accurately estimate the membrane resistance over a dendrite.

2. Formulation

In this section, we describe the three probabilistic models that we use in our method. Using these probabilistic models enables us to estimate the membrane-resistance over a dendrite from observed noisy membrane potential. In Section 2.1, we describe the cable equation [20], which expresses the dynamics of the den-

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dritic membrane potential, and its spatially discrete approximation, the compartment model [21]. We then derive the stationary distribution of the compartment model. In Section 2.2, we explain the smoothness prior, based on the MRF [18], [19], of the membrane resistance. The smoothness prior assumes that spatially adjacent membrane resistances take similar values, to accurately estimate membrane resistance over a dendrite, even when observation process is noisy. In Section 2.3, we describe the observation model, which expresses the noisy observation of membrane potential.

2.1 Cable Equation and Stationary Distribution for Compartment Model

In the cable equation [20], the dynamics of the membrane potential is given as

$$C \frac{\partial v(x, t)}{\partial t} = -a_x (v(x, t) - v_{\text{rev}}) + D \frac{\partial^2 v(x, t)}{\partial x^2} + u(x, t) + \sigma \xi(x, t), \quad (1)$$

where $v(x, t)$ is the membrane potential at position x at time t . In this paper, we consider a one-dimensional dendrite for the sake of simplicity. The right-hand side of Eq.(1) consists of four terms. The first term $-a_x (v(x, t) - v_{\text{rev}})$ expresses a passive linear membrane current, where a_x is the membrane conductance (inverse of membrane resistance) at position x and v_{rev} expresses reversal potential. The objective of our study was to estimate membrane conductance a_x from the observed membrane potential. The second term $D \frac{\partial^2 v(x, t)}{\partial x^2}$ expresses a current along the dendrite, where D is the intercompartmental conductance. The third term $u(x, t)$ expresses an external input, and the last term $\sigma \xi(x, t)$ expresses the internal noise of the neuron that is assumed to be white Gaussian with average $\langle \xi(x, t) \rangle = 0$ and correlation function $\langle \xi(x, t) \xi(x', t') \rangle = \delta(x - x') \delta(t - t')$. Parameter C on the left-hand side of Eq.(1) is the membrane capacitance. We can assume $C = 1$ without loss of generality. Next, we introduce a spatially discrete approximation to the cable equation: the compartment model [21]. A schematic of the compartment model is shown in Fig. 1. In this model, a dendrite is segmented into small compartments and the cable Eq. (1) is approximated as follows:

$$v_{x,t+1} - v_{x,t} = \Delta t \{-a_x (v_{x,t} - v_{\text{rev}}) + D(v_{x-1,t} - 2v_{x,t} + v_{x+1,t}) + u_{x,t}\} + \sqrt{\Delta t} \epsilon_{x,t}, \quad (2)$$

where $v_{x,t}$ and $u_{x,t}$ are the membrane potential and the external input, respectively. The last term $\sqrt{\Delta t} \epsilon_{x,t}$ is the internal noise assumed to be Gaussian with mean 0 and variance $\Delta t \sigma^2$ at compartment x at time t , and is derived by discretizing Langevin noise term $\sigma \xi(x, t)$ in the cable Eq. (1). Note that the factor $\sqrt{\Delta t}$ is important so that the noise variance grows linearly with time t .

We derive the stationary distribution of Eq.(2) for computational simplicity. Let $\tilde{v}_t = v_t - v_{\text{rev}}$ in Eq.(2), where v_t and v_{rev} are M -dimensional column vectors $(v_{1,t}, \dots, v_{M,t})^T$ and $(v_{\text{rev}}, \dots, v_{\text{rev}})^T$, respectively. M is the number of compartments. We then obtain

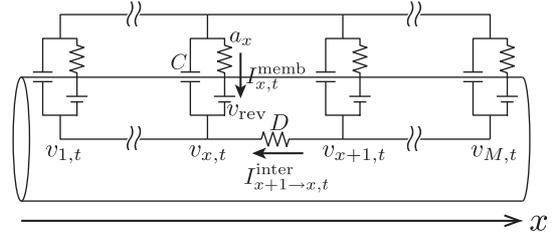


Fig. 1 Schematic of compartment model. Compartment model is a spatially discrete approximation of the cable equation. Membrane current $I_{x,t}^{\text{memb}}$ and current along dendrite $I_{x+1 \rightarrow x,t}^{\text{inter}}$ are given as $-a_x(v_{x,t} - v_{\text{rev}})$ and $D(v_{x+1,t} - v_{x,t})$, respectively. $v(x, t)$ is membrane potential at position x at time t , which expresses electrical state of dendrite. a_x , D , and C are membrane conductance at position x , intercompartmental conductance, and membrane capacitance, respectively. These three parameters define electrical property of dendrite. Stationary distribution of this compartment model is expressed as a Gaussian distribution. In our method, membrane conductance a_x is estimated using Markov Random Field (MRF). Unlike the previous studies, in our method membrane conductance a_x is not expressed as a function of position on dendrite, but is assumed to be smoothly varying along dendrite by using MRF.

$$\tilde{v}_{t+1} = \Phi \tilde{v}_t + \Delta t \left(\mathbf{u}_t + \frac{1}{\sqrt{\Delta t}} \boldsymbol{\epsilon}_t \right), \quad (3)$$

$$\Phi = I - \Delta t \Psi, \quad (4)$$

$$\Psi = \begin{pmatrix} a_1 & & & & \\ & \ddots & & & \\ & & a_M & & \\ & & & & \end{pmatrix} + D \begin{pmatrix} 1 & -1 & & & \\ -1 & 2 & -1 & & \\ & \ddots & \ddots & \ddots & \\ & & & -1 & 2 & -1 \\ & & & & -1 & 1 \end{pmatrix}, \quad (5)$$

where $\mathbf{u}_t = (u_{1,t}, \dots, u_{M,t})^T$, $\boldsymbol{\epsilon}_t = (\epsilon_{1,t}, \dots, \epsilon_{M,t})^T$, and I is the identity matrix. This equation is a first-order autoregressive model with Gaussian noise. If we keep the external input \mathbf{u}_t constant ($\mathbf{u}_t = \mathbf{u}$), the probability density function of the true membrane potential converges to the stationary distribution as $t \rightarrow \infty$. Since Eq. (3) is a Gaussian process, the stationary distribution is a Gaussian distribution. Therefore, we just need to determine the mean and covariance of the distribution. First, we derive the mean of the stationary distribution $E[\tilde{v}_\infty]$. By iteratively solving Eq. (3), we obtain

$$\tilde{v}_t = \Phi^t \tilde{v}_0 + \Delta t \sum_{s=0}^{t-1} \Phi^s \left(\mathbf{u} + \frac{1}{\sqrt{\Delta t}} \boldsymbol{\epsilon}_{t-1-s} \right). \quad (6)$$

Since $E[\boldsymbol{\epsilon}_t] = \mathbf{0}$,

$$E[\tilde{v}_\infty] = \Delta t (I - \Phi)^{-1} \mathbf{u} \approx \Psi^{-1} \mathbf{u}, \quad (7)$$

where we used $\lim_{t \rightarrow \infty} \Phi^t = 0$ and $\sum_{s=0}^{\infty} \Phi^s = (I - \Phi)^{-1}$. Next, we derive the covariance matrix $\text{Cov}[\tilde{v}_\infty]$. From Eq. (6),

$$\begin{aligned} \text{Cov}[\tilde{v}_t] &= \text{Cov} \left[\Phi^t \tilde{v}_0 + \Delta t \sum_{s=0}^{t-1} \Phi^s \left(\mathbf{u} + \frac{1}{\sqrt{\Delta t}} \boldsymbol{\epsilon}_{t-1-s} \right) \right] \\ &= \Delta t \sigma^2 \sum_{s=0}^{t-1} \Phi^{2s}. \end{aligned} \quad (8)$$

By taking the limit $t \rightarrow \infty$,

$$\text{Cov}[\tilde{v}_\infty] = \Delta t \sigma^2 (I - \Phi^2)^{-1} \approx \frac{\sigma^2}{2} \Psi^{-1}. \quad (9)$$

Thus, the stationary distribution is given as a Gaussian distribution:

$$p(\mathbf{v}|\mathbf{a}) = \mathcal{N}\left(\mathbf{v} \mid \mathbf{v}_{\text{rev}} + \Psi^{-1}\mathbf{u}, \frac{\sigma^2}{2}\Psi^{-1}\right). \quad (10)$$

We omit the subscript ∞ for the sake of notational simplicity. We can rewrite Eq. (10) using an energy function $E(\mathbf{v}|\mathbf{a})$:

$$p(\mathbf{v}|\mathbf{a}) = \frac{1}{Z(\mathbf{a})} \exp\left(-\frac{1}{\sigma^2}E(\mathbf{v}|\mathbf{a})\right), \quad (11)$$

$$E(\mathbf{v}|\mathbf{a}) = \sum_{x=1}^M a_x (v_x - \bar{v}_x)^2 + D \sum_{x=1}^{M-1} (v_{x+1} - v_x)^2, \quad (12)$$

$$Z(\mathbf{a}) = (\pi\sigma^2)^{\frac{M}{2}} |\Psi|^{-\frac{1}{2}}, \quad (13)$$

where \bar{v}_x is the x -th element of $\mathbf{v}_{\text{rev}} + \Psi^{-1}\mathbf{u}$.

2.2 Prior Distribution of Membrane Conductance

In this section, we introduce the smoothness prior, based on the MRF [18], [19], of the membrane conductance. The MRF is represented by a probability density function:

$$p(\mathbf{a}) \propto \exp(-E(\mathbf{a})), \quad (14)$$

$$E(\mathbf{a}) = \lambda \sum_{x=1}^{M-1} (a_{x+1} - a_x)^2, \quad (15)$$

$$a_x \in [0, \infty). \quad (16)$$

This equation expresses a physiological premise that membrane conductances of nearby compartments take similar values. The probability $p(\mathbf{a})$ increases if nearby membrane conductances take similar values and decreases if they take dissimilar ones. The factor λ is called hyperparameter. In this paper, we show the results in which λ was set to 100. We changed λ from 20 to 200 and obtained qualitatively similar results to those in which $\lambda = 100$ (data not shown). The performance of our method is thus robust to changes in λ .

As mentioned above, the objective of our study was to estimate the membrane conductance a_x over the dendrite. Accurate estimation of the membrane conductance has been difficult because the signal-to-noise ratio of membrane potential imaging is low. We use the MRF as a prior distribution of membrane conductance, to accurately estimate the membrane conductance over the dendrite even when observation process is noisy, without expressing the membrane conductance as a function of position on dendrite.

2.3 Observation Model

We introduce the observation model, a Gaussian process, which expresses the noisy observation of membrane potential. Let $\mathbf{y}_t = (y_{1,t}, \dots, y_{M,t})^T$ be the observed membrane potential at time t . Then, the observation model is given as

$$p(\mathbf{y}_t|\mathbf{v}_t) = \frac{1}{(2\pi\eta^2)^{\frac{M}{2}}} \exp\left(-\frac{1}{2\eta^2}E(\mathbf{y}_t|\mathbf{v}_t)\right), \quad (17)$$

$$E(\mathbf{y}_t|\mathbf{v}_t) = \sum_{x=1}^M (y_{x,t} - v_{x,t})^2. \quad (18)$$

This equation expresses that the observed membrane potential $y_{x,t}$

is the sum of the true membrane potential $v_{x,t}$ and Gaussian noise with variance η^2 .

3. Estimation

In this section, we illustrate the estimation method. By using the models Eqs. (11)–(18) described above, we estimate membrane conductance a_x and potential v_x from observed noisy data y_x . We derive the estimation method based on the EM algorithm [22]. The EM algorithm is a standard method for estimating parameters in statistical models based on the maximum likelihood or the maximum a posteriori principles.

The EM algorithm iterates over two steps, expectation (E-step) and maximization (M-step). In the E-step, we obtain the expectation value of the membrane potential \mathbf{v} , and in the M-step, we obtain the estimates of the membrane conductance \mathbf{a} . Let $\mathbf{Y} = \{\mathbf{y}_1, \dots, \mathbf{y}_N\}$ denote a set of observed membrane potentials and $\mathbf{V} = \{\mathbf{v}_1, \dots, \mathbf{v}_N\}$ denote a set of corresponding true membrane potentials. Then, the two steps are given as follows:

E-step Based on the current estimate of the parameter \mathbf{a}_{old} , the conditional distribution of the latent variables $p(\mathbf{V}|\mathbf{Y}, \mathbf{a}_{\text{old}})$ is calculated. Then the expected values of \mathbf{V} , and the expected complete-data loglikelihood $Q(\mathbf{a}, \mathbf{a}_{\text{old}}) = \langle \log p(\mathbf{Y}, \mathbf{V}|\mathbf{a}) \rangle_{p(\mathbf{V}|\mathbf{Y}, \mathbf{a}_{\text{old}})}$ are computed.

$$Q(\mathbf{a}, \mathbf{a}_{\text{old}}) = \frac{N}{2} \log |\Psi| - \frac{1}{\sigma^2} \sum_i^N \left\{ \text{Tr}(\Psi(\Sigma + \mathbf{v}_{\text{rev}}\mathbf{v}_{\text{rev}}^T)) + \mathbf{m}_i^T \Psi \mathbf{m}_i - 2\mathbf{v}_{\text{rev}} \mathbf{m}_i^T \mathbf{a} + \mathbf{u}^T \Psi^{-1} \mathbf{u} \right\} + \text{const.}, \quad (19)$$

where \mathbf{m}_i, Σ are the mean and the covariance of the Gaussian distribution $p(\mathbf{v}_i|\mathbf{y}_i, \mathbf{a}_{\text{old}})$. Equation (19) is derived in Appendix.

M-step A new estimation value of the parameter \mathbf{a}_{new} is inferred, which maximizes the sum of $Q(\mathbf{a}, \mathbf{a}_{\text{old}})$ and $\log p(\mathbf{a})$:

$$\mathbf{a}_{\text{new}} = \underset{\mathbf{a}}{\text{argmax}} \{Q(\mathbf{a}, \mathbf{a}_{\text{old}}) + \log p(\mathbf{a})\}. \quad (20)$$

Starting with the initial setting $\mathbf{a}_{\text{old}} = \mathbf{a}_0$, these two steps are repeated until convergence.

4. Results

We present results of applying our method to synthetic data. The synthetic data were generated as follows. First, true membrane potentials were generated from the compartment model, Eq. (2). Observed membrane potentials \mathbf{Y} were then generated from the observation model, Eq. (17). We estimated membrane conductance \mathbf{a} and membrane potentials \mathbf{V} from observed membrane potential \mathbf{Y} generated as above. We compared our method to that without the MRF, in which $p(\mathbf{a})$ is a uniform distribution instead of Eq. (14). We set $D = 10$, $v_{\text{rev}} = -70$, $\sigma = 0.01$, $\Delta t = 0.01$, $\eta = 0.05$, and $\lambda = 100$. The number of samples N was 200.

4.1 Sigmoidal Case

First, we present the results of applying the methods to the case where membrane conductance varies sigmoidally, plotted as

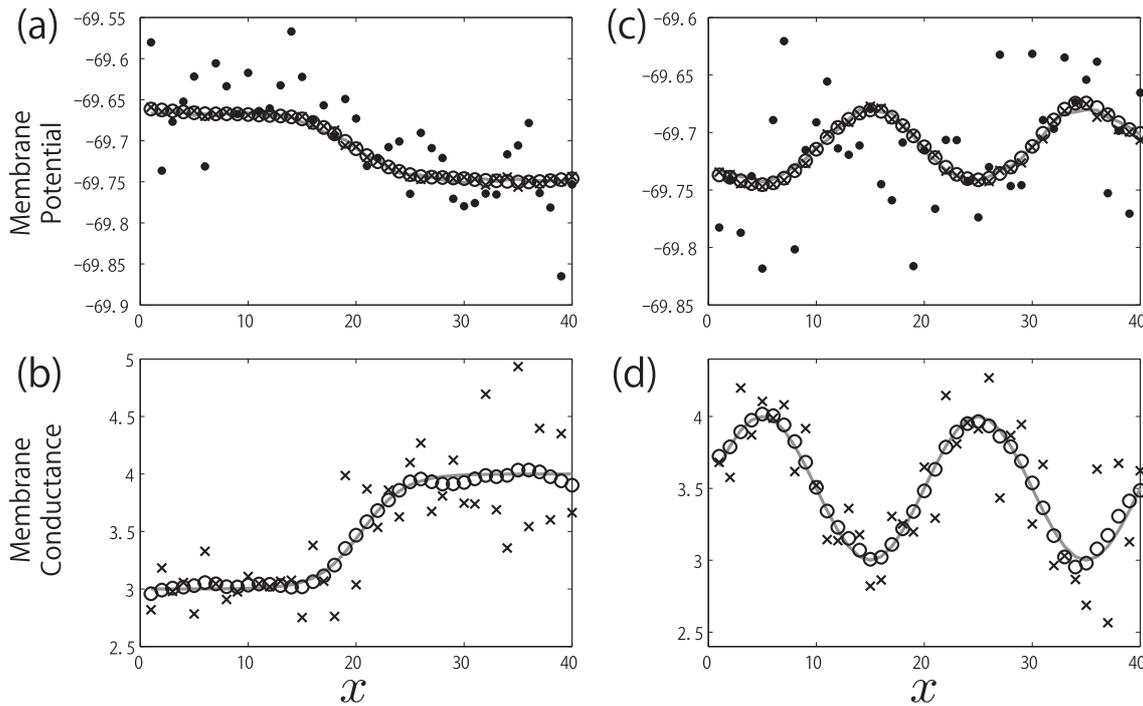


Fig. 2 Estimating parameters for sigmoidal case ((a), (b)) and sinusoidal case ((c), (d)). Top panels show membrane potential. Sample of observed membrane potential out of N samples is plotted as black circles. Corresponding true membrane potential, estimate using MRF, and estimate without MRF are plotted as gray line, open circles (\circ), and crosses (\times), respectively. Bottom panels show membrane conductance. True membrane conductance, estimate using MRF, and estimate without MRF are plotted as gray line, open circles (\circ), and crosses (\times), respectively.

a gray line in **Fig. 2**(b). In hippocampal CA1 pyramidal neuron dendrite, sigmoidally-varying membrane-conductance is observed [4], [5]. A sample of observed membrane potential out of N samples is plotted as black circles in Fig. 2(a). The corresponding true membrane potential, estimate using the MRF, and estimate without the MRF are plotted as gray line, open circles (\circ), and crosses (\times), respectively. We can see that the open circles (\circ) and crosses (\times) are almost on the gray line, that is, the estimates of membrane potential agree well with the true membrane potential. The estimates of membrane conductance are plotted in Fig. 2(b). Although our method did not assume that the membrane conductance varies sigmoidally, estimated membrane conductance (\circ) agrees well with the true membrane conductance. In contrast, the estimate without MRF (\times) is less accurate.

4.2 Sinusoidal Case

Second, we present results of applying the methods to the case where membrane conductance varies sinusoidally, to show that our method is applicable not only to the sigmoidal case. As is in the above case, observed membrane potential is plotted as black circles in Fig. 2(c). The true membrane potential, the estimate using the MRF, and the estimate without the MRF are plotted as gray line, open circles (\circ), and crosses (\times), respectively. The membrane potential is plotted in Fig. 2(c). We can see that the open circles (\circ) and crosses (\times) are almost on the gray line, that is, the estimates of membrane potential agree well with the true membrane potential. The estimates of membrane conductance is plotted in Fig. 2(d). The estimate using the MRF (\circ) agree well with the true membrane conductance, while the estimate without

the MRF (\times) deviates due to noise.

As presented above, in both sigmoidal and sinusoidal cases, the membrane conductances estimated using the MRF agree well with true membrane conductances, while those estimated without the MRF deviate from the true membrane conductances. Thus, our method, in which the MRF is used as a smoothness prior, enables us to estimate the membrane-resistance over the dendrite accurately even when the appropriate function is unknown.

5. Summary

We proposed a method for estimating the membrane resistance over a dendrite. The dynamics of the membrane potential are expressed using the compartment model and the observation process was modeled as a Gaussian process. Membrane resistance was estimated using the EM algorithm.

Unlike the previous studies, in our method membrane conductance is not expressed as a function of position on dendrite, but is assumed to be smoothly varying along dendrite by using the MRF. We showed using synthetic data that our method can be applied when the appropriate function is unknown.

The stationary distribution of the compartment model is used for computational simplicity. Transient dynamics can be used for estimation by applying Kalman filter to the compartment model. We targeted voltage-independent resistance. Our framework using MRF as a prior distribution of membrane resistance can also be applicable to voltage-dependent resistance. This is a subject for further study.

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Appendix

A.1 Derivation of the Expected Complete-data Loglikelihood

In our method we use the EM algorithm. In the M-step of the EM algorithm, the expected complete-data loglikelihood is computed. In this appendix, we give the derivation of the expected complete-data loglikelihood Eq. (19).

The conditional distribution $p(\mathbf{V}|\mathbf{Y}, \mathbf{a}_{\text{old}})$ is given by

$$\begin{aligned} p(\mathbf{V}|\mathbf{Y}, \mathbf{a}_{\text{old}}) &= \prod_{i=1}^N p(\mathbf{v}_i|\mathbf{y}_i, \mathbf{a}_{\text{old}}) \\ &= \prod_{i=1}^N \mathcal{N}(\mathbf{v}_i|\mathbf{m}_i, \Sigma), \end{aligned} \quad (\text{A.1})$$

where

$$\mathbf{m}_i = \Sigma \left\{ \eta^{-2} \mathbf{y}_i + 2\sigma^{-2}(\mathbf{u} + \Psi \mathbf{v}_{\text{rev}}) \right\}, \quad (\text{A.2})$$

$$\Sigma = \eta^{-2} + 2\sigma^{-2}\Psi. \quad (\text{A.3})$$

Then, the expected complete-data loglikelihood $Q(\mathbf{a}, \mathbf{a}_{\text{old}})$ is given by

$$\begin{aligned} Q(\mathbf{a}, \mathbf{a}_{\text{old}}) &= \langle \log p(\mathbf{Y}, \mathbf{V}|\mathbf{a}) \rangle_{p(\mathbf{V}|\mathbf{Y}, \mathbf{a}_{\text{old}})} \\ &= \sum_{i=1}^N \langle \log p(\mathbf{y}_i|\mathbf{v}_i) + \log p(\mathbf{v}_i|\mathbf{a}) \rangle_{p(\mathbf{v}_i|\mathbf{y}_i, \mathbf{a}_{\text{old}})} \\ &= \sum_{i=1}^N \left\langle -\frac{1}{2} \log \left| \frac{\sigma^2}{2} \Psi^{-1} \right| - \frac{1}{\sigma^2} (\mathbf{v}_i - \boldsymbol{\mu})^T \Psi (\mathbf{v}_i - \boldsymbol{\mu}) \right\rangle_{p(\mathbf{v}_i|\mathbf{y}_i, \mathbf{a}_{\text{old}})} \\ &\quad + \text{const.}, \end{aligned} \quad (\text{A.4})$$

where

$$\boldsymbol{\mu} = \mathbf{v}_{\text{rev}} + \Psi^{-1} \mathbf{u}. \quad (\text{A.5})$$

The distribution $p(\mathbf{v}_i|\mathbf{y}_i, \mathbf{a}_{\text{old}})$ is a Gaussian distribution with mean \mathbf{m}_i and covariance Σ . Hence, $Q(\mathbf{a}, \mathbf{a}_{\text{old}})$ is given by

$$\begin{aligned} Q(\mathbf{a}, \mathbf{a}_{\text{old}}) &= \frac{N}{2} \log |\Psi| - \frac{1}{\sigma^2} \sum_{i=1}^N \left\{ (\mathbf{m}_i - \boldsymbol{\mu})^T \Psi (\mathbf{m}_i - \boldsymbol{\mu}) + \text{Tr}(\Psi \Sigma) \right\} + \text{const.} \\ &= \frac{N}{2} \log |\Psi| - \frac{1}{\sigma^2} \sum_{i=1}^N \left\{ \text{Tr}(\Psi(\Sigma + \mathbf{v}_{\text{rev}} \mathbf{v}_{\text{rev}}^T)) + \mathbf{m}_i^T \Psi \mathbf{m}_i \right. \\ &\quad \left. - 2\mathbf{v}_{\text{rev}} \mathbf{m}_i^T \mathbf{a} + \mathbf{u}^T \Psi^{-1} \mathbf{u} \right\} + \text{const.}. \end{aligned} \quad (\text{A.6})$$



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